The Law of the Iterated Logarithm For a Stationary Process

> Michael Woodroofe The University of Michigan Joint work with Ou Zhao

# Limit Theorems

Let  $X_1, X_2, \cdots$  be independent and identically distributed random variables defined on a probability space  $(\Omega, \mathcal{A}, P)$ . Thus,

$$P[X_k \in B_k, k = 1, ..., n] = \prod_{k=1}^n P[X_1 \in B_k]$$

for Borel sets  $B_k \subseteq \mathbb{R}$ . Suppose that  $X_k$  have a finite mean and (not necessarily finite) variance

$$E(X_k) = \mu$$
 and  $\sigma^2 = E[(X_k - \mu)^2],$ 

and let

$$S_n = X_1 + \dots + X_n.$$

# Outline

Classical Limit Theorems.

Stationary Processes.

The Conditional Central Limit Theorem.

The Law of the Iterated Logarithm.

The Functional Version.

The Proof.

Questions.

The Law of Large Numbers.

$$\lim_{n \to \infty} \frac{S_n}{n} = \mu \ w.p.1.$$

The Central Limit Theorem. If  $0 < \sigma^2 < \infty$ , then

$$\lim_{n \to \infty} P\left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \le z\right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-\frac{1}{2}x^2} dx = \Phi(z),$$

The Law of the Iterated Logarithm. If  $\sigma^2 < \infty$ , then

$$\limsup_{n \to \infty} \frac{S_n - n\mu}{\sqrt{2n \log \log(n)}} = \sigma \ w.p.1.$$

**Note**: Physical interpretation of  $\mu$  and  $\sigma$ .

**Corollary** For large n,

$$P\left[|\frac{S_n - n\mu}{\sigma\sqrt{n}}| \le 3\right] \approx \frac{1}{\sqrt{2\pi}} \int_{-3}^3 e^{-\frac{1}{2}x^2} dx = .9974 \approx 1.$$

**Corollary**. If  $0 < \sigma^2 < \infty$ , then

$$P\left[\limsup_{n\to\infty}\left|\frac{S_n-n\mu}{\sigma\sqrt{n}}\right|>3\right]=1.$$

Remarks

- A single n versus for some n
- Statistical Paradox

### A Statistical Paradox

**Recall**:  $X_1, X_2, \cdots$  are i.i.d. with mean  $\mu$  and variance  $\sigma^2$ .

An Hypothesis:  $\mu = \mu_0$  and  $\sigma = \sigma_0$ .

A Test: Reject if

$$\left|\frac{S_n - n\mu_0}{\sigma_0\sqrt{n}}\right| > 3.$$

Then

$$P_0[\text{Reject}] < .01.$$

**Optional Stopping.** Now let N be the smallest  $n \ge 1$  for which  $|S_n - n\mu_0| > 3\sigma_0\sqrt{n}$ . Then

 $N < \infty w.p.1$  and  $P_0[\text{Reject}] = 1.$ 

Example: ESP

Question: To Bayes or not to Bayes.

# **Stationary Processes**

**Def.** A sequence  $\cdots X_{-1}, X_0, X_1, \cdots$  is stationary iff

 $P\left[\left(X_{n+1},\cdots,X_{n+m}\right)\in B\right]$ 

is independent of n for all Borel sets  $B \subseteq \mathbb{R}^m$  for all m.

Sequence Space: Let  $\mathbf{X} = (\cdots, X_{-1}, X_0, X_1, \cdots)$ , and

$$Q(B) = P[\mathbf{X} \in B]$$

for Borel sets  $B \subseteq \mathbb{R}^Z$ ; and let  $\theta$  be the shift operator

$$\theta(\cdots x_{-1}, x_0, x_1, \cdots) = (\cdots, x_0, x_1, x_2, \cdots).$$

Then the sequence is stationary iff  $Q \circ \theta^{-1} = Q$ .

**Ergodicity**: A stationary sequence is ergodic if Q(A) = 0 or 1 whenever  $\theta^{-1}(A) = A$ .

#### **Examples**

Let  $\cdots X_{-1}, X_0, X_1, \cdots$  be independent and identically distributed; and let  $\psi : \mathbb{R}^Z \to \mathbb{R}$  be measurable. Then

$$Y_k = \psi(\cdots X_{k-1}, X_k, X_{k+1}, \cdots)$$

is stationary and ergodic.

### Martingale Differences

Let  $\cdots X_{-1}, X_0, X_1, \cdots$  be a sequence of random variables and  $\mathcal{F}_k = \sigma \{\cdots X_{k-1}, X_k\}$ . The random variables are martingale differences if

 $E|X_k| < \infty$  and  $E(X_k|\mathcal{F}_{k-1}) = 0 \ w.p.1$ 

for all k in which case  $M_n = X_1 + \cdots + X_n$  is called a martingale.

**Example**: If  $\cdots X_{-1}, X_0, X_1, \cdots$  are independent and identically distributed (IID) with  $E(X_k) = 0$ , then they form martingale differences.

The CLT and LIL: The CLT and LIL are valid for stationary sequences of martingale differences in which case  $\mu = 0$ .

Notation:  $||Y|| = \sqrt{E(Y^2)}$ .

About the Condition:  $\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \| E(S_n | \mathcal{F}_0) \| < \infty$ .

- Limits the dependence.
- Solely in terms of  $||E(S_n|\mathcal{F}_0)||$ .
- Best Possible.
- Compare with strong mixing.

Examples: Bernoulli shifts; Hasting Metroplis; linear proceses.

Refs: Maxwell and W (2000), Ann. Prob.

Peligrad and Utev (2005) Ann. Prob.

# A CCLT For Stationary Sequences

Let  $\cdots X_{-1}, X_0, X_1, \cdots$  be a statonary, ergodic sequence with  $E(X_k) = 0$  and  $E(X_k^2) < \infty$ ; let  $\mathcal{F}_k = \sigma\{\cdots X_{k-1}, X_k\}$ ; and let  $S_n = X_1 + \cdots + X_n$ . If

$$\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \|E(S_n|\mathcal{F}_0)\| < \infty, \qquad (\dagger)$$

then

$$\sigma^2 = \lim_{n \to \infty} \frac{1}{n} E(S_n^2) \tag{(*)}$$

exists (finite), and

$$\lim_{n \to \infty} P\left[\frac{S_n}{\sqrt{n}} \le z |\mathcal{F}_0\right] = \Phi(\frac{z}{\sigma}) \tag{**}$$

in probability for all z. Conversely if (\*) and (\*\*), then

$$||E(S_n|\mathcal{F}_0)|| = o(\sqrt{n}).$$

#### About the Proof

In the proof it is shown that

$$S_n = M_n + R_n$$

where  $M_n$  and  $R_n$  are  $\mathcal{F}_n$ -measurable,  $M_n$  is the sum of a stationary sequence of martingale differences, and  $||R_n|| = o(\sqrt{n})$ .

Connection to LIL: If

$$R_n = o[\sqrt{n \log \log(n)}] \ w.p.1,$$

then the LIL would hold for  $S_n$ .

## **Slow Variation**

A function  $\ell: (0,\infty) \to (0,\infty)$  varies slowly at  $\infty$  iff

$$\lim_{t \to \infty} \frac{\ell(tx)}{\ell(t)} = 1$$

for all  $0 < x < \infty$ .

Notation: Let

$$\ell^*(n) = \sum_{j=1}^n \frac{1}{j\ell(j)}$$

**Example:** If  $\ell(n) = \log(n+1)$ , then  $\ell^*(n) \sim \log \log(n)$ .

# A Law of the Iterated Logarithm

Let  $\cdots X_{-1}, X_0, X_1, \cdots$  be a stationary, ergodic sequence with  $E(X_k) = 0$  and  $E(X_k^2) < \infty$ ; let  $\mathcal{F}_k = \sigma\{\cdots X_{k-1}, X_k\}$ ; and let  $S_n = X_1 + \cdots + X_n$ . If  $\ell : (0, \infty) \to (0, \infty)$  is a non-decreasing, slowly varying function, for which

$$\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \log(n) \sqrt{\ell(n)} \| E(S_n | \mathcal{F}_0) \| < \infty, \qquad (\dagger)$$

then

$$\lim_{n \to \infty} \frac{R_n}{\sqrt{n\ell^*(n)}} = 0 \ w.p.1.$$

Corollary 2. If

$$\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \log(n) \sqrt{\ell(n)} \| E(S_n | \mathcal{F}_0) \| < \infty, \qquad (\dagger)$$

for some  $\ell$  for which  $1/[n\ell(n)]$  is summable, then

$$\lim_{n \to \infty} \frac{R_n}{\sqrt{n}} = 0 \ w.p.1,$$

and there is convergence w.p.1 in the Conditional Central Limit Theorem. That is,

$$\lim_{n \to \infty} P\left[\frac{S_n}{\sqrt{n}} \le z |\mathcal{F}_0\right] = \Phi(\frac{z}{\sigma}) \ w.p.1. \tag{*}$$

**Corollary 1.** If  $(\dagger)$  holds with  $\ell(n) = \log(n+1)$ ,

$$\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \log^{\frac{3}{2}}(n) \| E(S_n | \mathcal{F}_0) \| < \infty, \tag{\dagger}$$

then

$$\lim_{n \to \infty} \frac{R_n}{\sqrt{n \log \log(n)}} = 0 \ w.p.1$$

and

$$\limsup_{n \to \infty} \frac{S_n}{\sqrt{2n \log \log(n)}} = \sigma \ w.p.1.$$

### The Functional Version

**Brownian Motion**. Let  $\mathbb{B}$  denote standard Brownian motion. Thus,  $\mathbb{B}(t)$ ,  $0 \le t \le 1$ , is a stochastic process with

- independent increments;
- continuous sample paths
- $I\!B(0) = 0$

•  $I\!\!B(t) - I\!\!B(s)$  is normally distributed with mean 0 and variance t - s.

**Sample Versions.** Let  $X_1, X_2, \cdots$  be random variables,  $S_n = X_1 + \cdots + X_n$ , and let  $I\!B_n$  be a continuous piecewise linear function for which

$$\mathbb{B}_n(\frac{k}{n}) = \frac{1}{\sqrt{n}} S_k, \ k = 0, 1, \cdots, n.$$

**Corollary 3.** Let  $\cdots X_{-1}, X_0, X_1, \cdots$  is be stationary process with mean 0 and finite variance. If

$$\sum_{n=1}^{\infty} n^{-\frac{3}{2}} \log^{\frac{3}{2}}(n) \| E(S_n | \mathcal{F}_0) \| < \infty, \tag{\dagger}$$

then w.p.1 the set of limit points of  $\mathbb{B}_n/\sqrt{2n\log\log(n)}$ ,  $n \ge 3$ , is  $\sigma K$ , where

$$\sigma^2 = \lim_{n \to \infty} \frac{1}{n} E(S_n^2).$$

*Proof.* Follow from Strassen's Theorem (for martingale differences) and  $R_n = o[\sqrt{2n \log \log(n)}]$ .

#### Strassen's LIL

Let C[0,1] be the set of continuous functions  $f:[0,1] \to \mathbb{R}$ , and let  $K \subseteq C[0,1]$  be the set of absolutely continuous f for which

$$\int_0^1 f'(t)^2 dt \le 1$$

If  $X_1, X_2, \cdots$  are i.i.d. with mean 0 and variance one, then the set of limit points of

$$\frac{I\!B_n}{\sqrt{2n\log\log(n)}}, n \ge 3$$

is K w.p.1.

# About the Proof Markov Chains

There is no loss of generality in supposing that

$$X_k = g(W_k),$$

where  $\cdots W_{-1}, W_0, W_1 \cdots$  is a stationary Markov chain with values in a measurable space  $\mathcal{W}$ , for we may let

$$W_k = (\cdots, X_{k-1}, X_k).$$

Let  $\pi$  and Q denote the stationary distribution and transition function of the chain,

$$\pi\{B\} = P[W_k \in B]$$

and

$$Q(w;B) = P[W_{n+1} \in B | W_n = w].$$

$$Qf(w) = \int_{\mathcal{W}} f(z)Q(w;dz) = E[f(W_{n+1})|W_n = w] \ a.e.$$

**Some Operators** 

for  $f \in L^1(\pi)$ . Then,

$$E(X_k|W_0) = Q^k g(W_0)$$

and

Let

$$E(S_n|W_1) = \sum_{k=1}^n Q^{k-1}g(W_1) = V_ng(W_1).$$

The Operator Norm. If  $S: L^2(P) \to L^2(P)$  is a continuous linear transformation, let

$$||S||_{op} = \sup_{||Y|| \le 1} ||SY||.$$

**The Operator** B(T). Since  $\beta_k$  are summable,

$$B(T) = \sum_{k=1}^{\infty} \beta_k T^k$$

converges in the operator norm. Let

$$A(z) = \frac{1}{1 - B(z)} = \sum_{k=0}^{\infty} \alpha_k z^k$$

Then A is continuous in  $|z| \leq 1 \neq z$ . A major problem is to make sense of A(T).

# **More Operators**

Next, let  $\theta$  denote the shift operator,  $W_n \circ \theta = W_{n+1}$  and

$$Tf = f \circ \theta.$$

Next, let

$$\beta_k = \frac{c}{k} \sum_{n=k}^{\infty} \frac{1}{\sqrt{n^3 \ell(n)}} \sim \frac{2c}{\sqrt{k^3 \ell(k)}}$$

where c is so chosen that  $\beta_1 + \beta_2 + \cdots = 1$ , and

$$B(z) = \sum_{k=1}^{\infty} \beta_k z^k$$

Then B(z) is continuous in  $|z| \leq 1$  and analytic in |z| < 1.

# About the Martingale

Recall that  $X_k = g(W_k)$ , where  $g \in L^2(\pi)$  and  $\int_{\mathcal{W}} g d\pi = 0$ , and let

$$h_{\epsilon} = \sum_{k=1}^{\infty} \frac{Q^{k-1}g}{(1+\epsilon)^k}$$

and  $H_{\epsilon}(w_0, w_1) = h_{\epsilon}(w_1) - Qh_{\epsilon}(w_0)$ . Then, from MW(2000),

$$H = \lim_{\epsilon \to 0} H_{\epsilon}$$

exists in an  $L^2$  space,  $M_n = H(W_0, W_1) + \dots + H(W_{n-1}, W_n)$ , and

$$R_n = \sum_{k=1}^n \xi_0 \circ \theta^k.$$

with  $\xi_0 = g(W_0) - H(W_{-1}, W_0).$ 

#### **Some Details**

About the Proof

First Step: Show  $\xi_0 \in [I - B(T)]L^2(P)$ ; that is  $\xi_0 = \eta_0 - B(T)\eta_0$ for some  $\eta_0 \in L^2(P)$ .

- Bound  $||R_n||$ .
- Fourier analysis of  $B(e^{it})$ .

Second Step: Show that

$$\lim_{n \to \infty} \frac{1}{\sqrt{n\ell^*(n)}} \sum_{k=1}^n \xi \circ \theta^k = 0 \ w.p.1$$

for any  $\xi \in [I - B(T)]L^2(P)$ . Uses a version of the Dominated Ergodic Theorem.

**Second Step.** Let  $Y_k$  be stationary and

$$Y^* = \sup_{n \ge 1} |\frac{Y_1 + \dots + Y_n}{n}|$$

Then the Dominated Ergodic Theorem states:

$$||Y^*||_p \le \left(\frac{p}{p-1}\right) ||Y_1||_p$$

for p > 1. Also

$$E[\sqrt{Y^*}] \le 2E[\sqrt{|Y_1|}].$$

First Step. Recall

$$A(z) = \frac{1}{1 - B(z)} = \sum_{k=0}^{\infty} \alpha_k z^k.$$

To make sense of  $A(T)\xi_0$ , we need to bound  $\alpha_k$ . From  $\beta_k \sim 2c/\sqrt{k^3\ell(k)}$ , we get

$$B(e^{it}) \sim \kappa_0 \sqrt{\frac{t}{\ell(t)}}$$

as  $t \to 0$ . Thus

and

$$\alpha_k - \alpha_{k+1} = O[\sqrt{\frac{\ell(n)}{n^3}}]$$

 $A(e^{it}) \sim \frac{\sqrt{\ell(t)}}{\kappa_0 \sqrt{t}}$ 

# **Questions:** The Iceberg

To what extent does the theory of random walks extend to stationary, ergodic sequences?

- Law of Large Numbers: Yes
- Central Limit Theorem:  $S_n = M_n + R_n$ .
- Law of the Iterated Logarithm:  $S_n = M_n + R_n$ .
- When is  $S_n = M_n + R_n$ ?
- Local Limit Theorems, etc. ...: Bits and pieces,
- Renewal Theorem: Partially in Lalley (1985, PTRF).
- Spitzer's Identity for Ladder Heights: Operator versions.